

No Association Between Alpha-1-Antichymotrypsin Polymorphism and Alzheimer's Disease in Koreans

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To examine the possible involvement of the alpha-1-antichymotrypsin gene (*ACT*) polymorphism in the manifestation of Alzheimer's disease (AD), we analyzed genotypes of the *ACT* and apolipoprotein E gene (*APOE*) among 110 Korean patients with probable AD and 209 nondemented controls. No significant difference was obtained in genotypic ($\chi^2=1.98$, $df=2$, $P>0.1$) and allelic frequencies ($\chi^2=1.61$, $df=1$, $P>0.1$) of *ACT* between the AD and control groups. No overexpression of the *ACT* A/A genotype and *ACT* A allele was found when we analyzed the late-onset AD patients and the early-onset AD patients, separately. Then we stratified the *ACT* genotypes based on the presence or absence of the *APOE* $\epsilon 4$ allele to evaluate the possible interaction between them. In the *APOE* $\epsilon 4$ -negative subjects, although the *ACT* A allele tended to be overexpressed in the AD group, the differences in the frequencies of the *ACT* A allele ($\chi^2=2.79$, $df=1$, $P>0.1$) and *ACT* A/A genotype ($\chi^2=0.16$, $df=1$, $P>0.1$) were not statistically significant. No significant overrepresentations of the *ACT* A allele ($\chi^2=0.02$, $df=1$, $P>0.1$) and *ACT* A/A genotype ($\chi^2=0.17$, $df=1$, $P>0.1$) were found in the *APOE* $\epsilon 4$ -positive subjects, either. In addition, the status of the *ACT* genotype did not influence the age-at-onset of AD ($F=0.03$, $df=2$, $P>0.1$).

Therefore, the *ACT* polymorphism does not contribute to the development of AD independently or interactively with the *APOE* $\epsilon 4$ allele in Koreans. *Am. J. Med. Genet.* 91:355–358, 2000. © 2000 Wiley-Liss, Inc.

KEY WORDS: Alzheimer's disease; alpha-1-antichymotrypsin; apolipoprotein E; allele frequency; polymorphism; risk factor; age-at-onset

INTRODUCTION

Although the apolipoprotein E gene (*APOE*) $\epsilon 4$ allele is a well-established genetic risk factor for Alzheimer's disease (AD), it is neither necessary nor sufficient to cause AD [Farrer et al., 1997], and thus additional genetic or nongenetic factors might be involved in the manifestation of AD. Alpha-1-antichymotrypsin (*ACT*) needs to be tested as such an additional genetic risk factor for AD, since it has been reported not only to play an important role in the pathogenesis of AD, but also to bind to the position 1-12 of amyloid beta protein, which is adjacent to the *APOE* binding position 12-28 [Abraham et al., 1988; Ma et al., 1994].

Kamboh et al. [1995] reported that the A/A genotype of the *ACT* gene (*ACT*) increased the *APOE* $\epsilon 4$ -conferred risk for AD twofold in heterozygotes for the *APOE* $\epsilon 4$ allele and threefold in homozygotes for the allele beyond the risk conferred to the *APOE* $\epsilon 4$ allele alone, suggesting that the *ACT* polymorphism might be a probable additional genetic risk factor for AD. However, in a series of successive studies, an association between the *ACT* polymorphism and AD has not been consistently replicated [Didierjean et al., 1997; Haines et al., 1996; Helisalmi et al., 1997; Itabashi et al., 1998; Muller et al., 1996; Nacmias et al., 1996; Yamanaka et al., 1998; Yoshizawa et al., 1997]. The allele frequencies of the *ACT* polymorphism differed among ethnic

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groups: the frequencies of the *ACT* A allele and A/A genotype were significantly higher in Nigerian Blacks and lower in the Japanese than those in the Caucasians [Kamboh et al., 1997; Yoshizawa et al., 1997].

The purpose of this study is to know whether the *ACT* polymorphism contributes to the development of AD and whether it influences on the age-at-onset of AD in Koreans.

MATERIALS AND METHODS

All the AD patients and nondemented controls used in the present study were unrelated Koreans. We administered the standardized CERAD Clinical Assessment Battery [Morris et al., 1989] and the modified Hachinski Ischemic Score (MHIS) [Hachinski et al., 1974]. Following this evaluation, a consensus committee meeting was held to diagnose each subject. Diagnosis for dementia was made according to *Diagnostic and Statistical Manual of Mental Disorders* (DSM-IV) criteria [American Psychiatric Association, 1994] and that for AD according to the NINCDS-ADRDA criteria [McKhann et al., 1984]. The subjects who were diagnosed as cognitively normal and got 4 or fewer on the MHIS were included in the normal control group. Informed consent was obtained for each subject.

Genomic DNA was extracted from peripheral blood leukocytes from the subjects. Genotyping for the *APOE* gene was done according to Wenham et al. [1991], and that for *ACT* according to Kamboh et al. [1995]. Allele and genotype frequencies of *ACT* and *APOE* were estimated by counting alleles and genotypes and calculating sample proportions. Hardy-Weinberg equilibrium was tested by the likelihood-ratio test [Lynch and Walsh, 1998]. Comparisons of the allele and genotype frequencies were made using chi-square analysis and Fisher's exact test. Logistic regression analysis was done to calculate the odds ratios (ORs) for AD. Comparison of the age-at-onset among groups with different genotypes were made using analysis of variance (ANOVA).

RESULTS

A total of 110 patients with probable AD and 209 controls were enrolled in the present study. Although the AD group was slightly older than the control group (70.0±7.8 years versus 68.2±4.6 years, $P < 0.05$), the two groups did not differ by gender. The AD group was comprised of 62 patients with late-onset AD (LOAD) and 48 patients with early-onset AD (EOAD). The mean age-at-onset of the EOAD patients was 58.7±4.3 years and that of the LOAD patients 72.7±5.0 years, while the mean age of the EOAD patients was 62.8±4.4 years and that of the LOAD patients 75.7±4.4 years. The EOAD and LOAD groups did not differ by gender nor degree of impairments in cognitive functions and activities in daily life (data not shown).

The distributions of the *ACT* genotypes and alleles of the AD, EOAD, LOAD, and control groups are presented in Table I. No significant difference in *ACT* genotype frequency ($\chi^2 = 1.98$, $df = 2$, $P > 0.1$) and its allele frequency ($\chi^2 = 1.61$, $df = 1$, $P > 0.1$) was obtained be-

TABLE I. Distribution of *ACT* Genotypes and Alleles Among Alzheimer's Disease (AD), Early-Onset AD Patients (EOAD), Late-Onset AD Patients (LOAD), and Control Groups*

<i>ACT</i>	Number (%) of individuals			
	AD with			Control
	All	EOAD	LOAD	
Genotype				
A/A	17 (15.5)	6 (12.5)	11 (17.7)	27 (12.9)
A/T	60 (54.5)	25 (52.1)	35 (56.5)	103 (49.3)
T/T	33 (30.0)	17 (35.4)	16 (25.8)	79 (37.8)
Total	110	48	62	209
Allele				
A	94 (42.7)	37 (38.5)	57 (46.0)	157 (37.6)
T	126 (57.3)	59 (61.5)	67 (54.0)	261 (62.4)
Total	220	96	124	418

*EOAD, AD developed before age 65; LOAD, AD developed at age 65 or later. Distribution of *ACT* genotypes in each group is in Hardy-Weinberg equilibrium.

tween the AD control groups. No overexpression of the *ACT* A/A genotype or *ACT* A allele was found when we analyzed the LOAD and EOAD patients separately. Then we stratified the *ACT* genotypes based on the presence or absence of the *APOE* $\epsilon 4$ allele to evaluate the possible interaction between them (Table II). In the *APOE* $\epsilon 4$ -negative subjects, although the *ACT* A allele tended to be overexpressed in the AD group, the differences in the frequencies of the *ACT* A allele ($\chi^2 = 2.79$, $df = 1$, $P > 0.1$) and *ACT* A/A genotype ($\chi^2 = 0.16$, $df = 1$, $P > 0.1$) were not statistically significant. No significant overrepresentation of the *ACT* A allele ($\chi^2 = 0.02$, $df = 1$, $P > 0.1$) and *ACT* A/A genotype ($\chi^2 = 0.17$, $df = 1$, $P > 0.1$) was found in the *APOE* $\epsilon 4$ -positive subjects, either. We also performed logistic regression analysis to examine the possible interaction between the *APOE* $\epsilon 4$ allele and *ACT* genotype. The occurrence of the *APOE* $\epsilon 4$ allele and the *ACT* genotype was entered in the first step to examine their independent effects and then the interaction between the two was entered in the second step to examine specifically for interactive effects (Table III). In the first step, the effect of the *APOE* $\epsilon 4$ allele was confirmed, whereas no effect of the *ACT* genotype was seen. Also, no significant additional effect was seen in the second step ($\chi^2 = 2.80$, $df = 2$, $P > 0.1$), i.e., the interaction between the occurrence of the *APOE* $\epsilon 4$ allele and the *ACT* genotype did not meet the

TABLE II. Distribution of the *ACT* Genotypes and Alleles Stratified by the Occurrence of the *APOE* $\epsilon 4$ Allele

<i>ACT</i> ^a	Number (%) of individuals with			
	<i>APOE</i> $\epsilon 4$ (-)		<i>APOE</i> $\epsilon 4$ (+)	
	AD ^b	Control	AD ^b	Control
Genotype				
A/A	10 (14.9)	24 (13.0)	7 (16.3)	3 (12.5)
A/T	41 (61.2)	90 (48.6)	19 (44.2)	13 (54.2)
T/T	16 (23.9)	71 (38.4)	17 (39.5)	8 (33.3)
Total	67	185	43	24
Allele				
A	61 (45.5)	138 (37.3)	33 (38.4)	19 (39.6)
T	73 (54.5)	232 (62.7)	53 (61.6)	29 (60.4)
Total	134	370	86	48

^a*ACT* genotypes in each group were in Hardy-Weinberg equilibrium.

^bAD = Alzheimer's disease.

TABLE III. Odds Ratio (OR) With 95% Confidence Interval (CI) for Developing Alzheimer's Disease for the Occurrence of the APOE ϵ 4 Allele and ACT Genotype*

	Independent model		Interactive model	
	OR	95% CI	OR	95% CI
APOE ϵ 4 allele	1.0		1.0	
Absence				
Presence	5.08	2.85–9.06	9.43	3.47–25.63
ACT genotype	1.0		1.0	
T/T				
A/T	1.538	0.88–2.65	2.02	1.05–3.90
A/A	1.58	0.73–3.42	1.85	0.74–4.62
APOE+/ACT interaction				
Absence	T/T		1.0	
Presence	A/T		0.34	0.09–1.22
Presence	A/A		0.59	0.09–3.73
Likelihood	376.65, df = 3		373.85, df = 5	

*No significant difference ($\chi^2 = 2.80$, $df = 2$, $P > 0.1$).

criteria for the entry into the regression equation ($W = 2.76$, $df = 2$, $P > 0.1$).

The ACT genotype did not influence on the age-at-onset of AD ($F = 0.03$, $df = 2$, $P > 0.1$ by ANOVA). The mean age-at-onset of the AD patients with A/A, A/T, and T/T genotypes were 65.9 ± 9.3 , 65.6 ± 15.3 , and 65.0 ± 7.8 years, respectively. Also, the influence of the ACT genotype on age-at-onset did not show a significant interaction with the occurrence of the APOE ϵ 4 allele ($F = 0.43$, $df = 2$, $P > 0.1$) or APOE ϵ 2 allele ($F = 0.93$, $df = 2$, $P > 0.1$).

DISCUSSION

Kamboh et al. [1995] reported that the ACT polymorphism conferred a significant risk for AD as well as modified the APOE ϵ 4-conferred risk for AD. However, in the present study, we could not find any independent association between AD and the ACT polymorphism, or a significant interaction between the ACT polymorphism and the APOE ϵ 4 allele in the development of AD. There are several possible reasons, including the ethnicity, type of AD, or other unknown differences in the populations, for the conflicting results. A difference in the allele frequency regarding the ACT polymorphism was evident between Koreans and Caucasians. The frequency of the ACT A allele in normal Koreans was significantly lower than that in Caucasian controls (37.6% versus 48.4%, $P < 0.01$) [Kamboh et al., 1997]. However, it is less likely that the difference of the frequency among ethnic groups is a main source of the conflicting results, since the effect of the ACT polymorphism was not consistent even within a single ethnic group, e.g., in the Japanese [Itabashi et al., 1998; Muramatsu et al., 1996; Yamanaka et al., 1998; Yoshiiwa et al., 1997; Yoshizawa et al., 1997]. Considering that the significant association between the ACT A/A genotype and AD was observed only in late-onset familial AD cases by a few studies [Nacmias et al., 1998; Yoshizawa et al., 1997], the difference in the distribution of age-at-onset may also have produced the conflicting results. In the present study, the ACT A allele and A/A genotype tended to be more prevalent in the

LOAD group than in the EOAD group, but the difference did not reach statistical significance (46.0% versus 38.5%, $P > 0.1$).

Talbot et al. [1996] reported that the AD patients with the ACT A/A genotype showed onset of AD six years earlier than those with the ACT A/T or T/T genotype in the APOE ϵ 4 positive group. However, we could not find any association between the number of the ACT A alleles and age-at-onset of AD either independently or interactively with the APOE ϵ 4 allele or ϵ 2 allele. Our results are in agreement with those by Nacmias et al. [1996] and Yoshizawa et al. [1997]. This discrepancy may be attributable to the difference in the population studied, e.g., the proportion of familial cases or other unknown factors contributing to the progression of AD other than the status of the APOE polymorphism. Actually, in Talbot et al. [1996], the ACT A/A genotype was not overexpressed in AD patients, and an interaction between the ACT A/A genotype and APOE ϵ 4 allele for the development of AD was not significant, either.

In conclusion, the ACT polymorphism does not significantly contribute to the development of AD independently or interactively with the APOE ϵ 4 allele in the Korean.

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